

Bayesian Estimation of Trends in Population-Level Health Metrics Using Disparate Data Sources

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MATHEMATICA
Policy Research

WSS Seminar

November 12, 2014

Bayesian
Estimation of
Health Metric
Trends

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Finucane

Motivation

'Shrinkage'

Blood
Pressure

Childhood
Undernutrition

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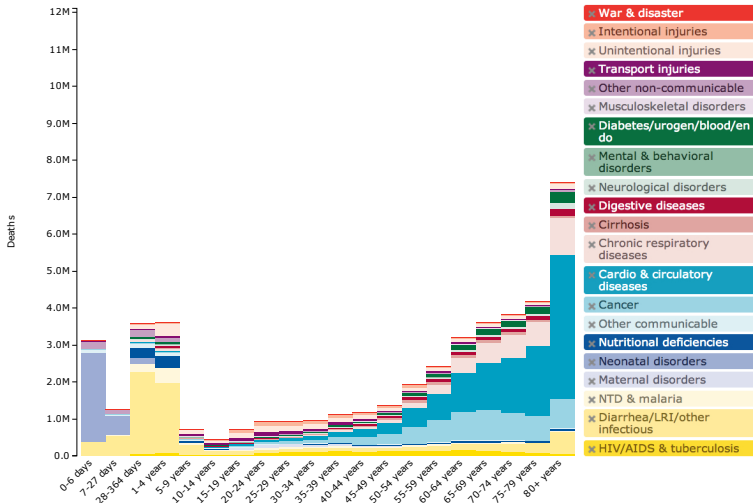
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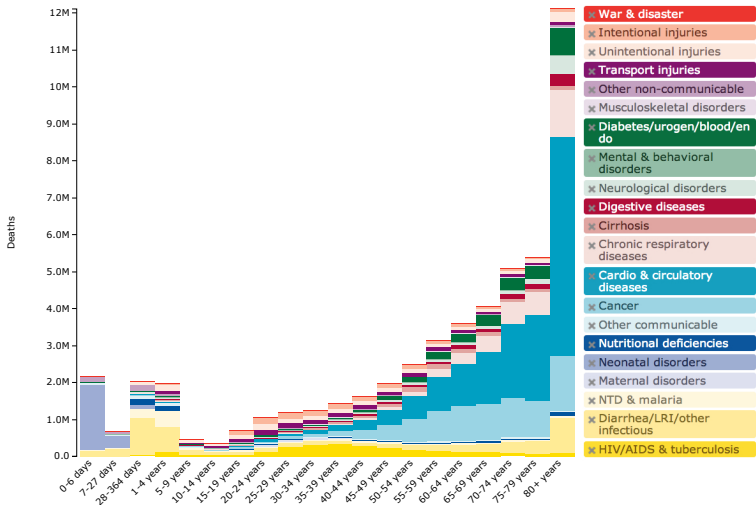
Childhood
Undernutrition



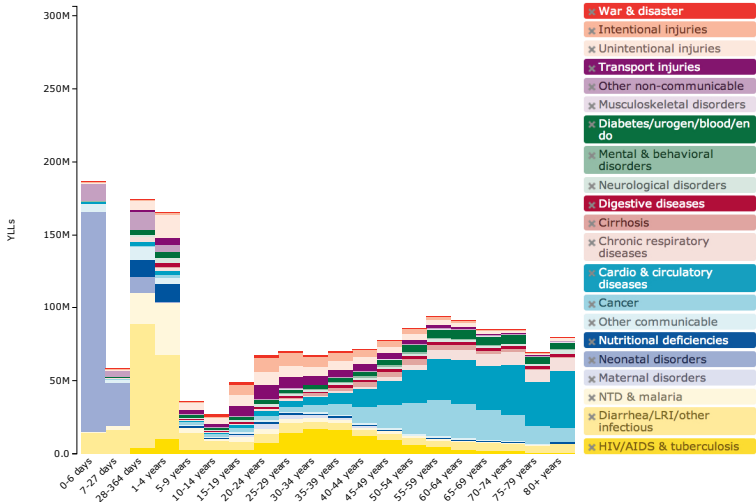
Deaths in 1990



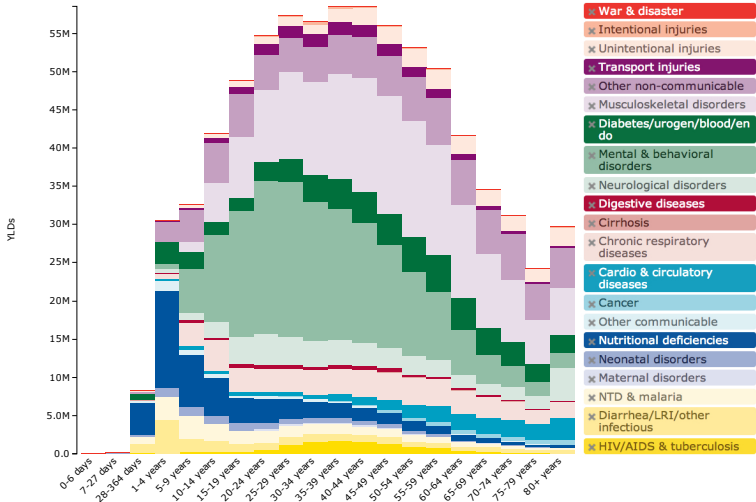
Deaths in 2010



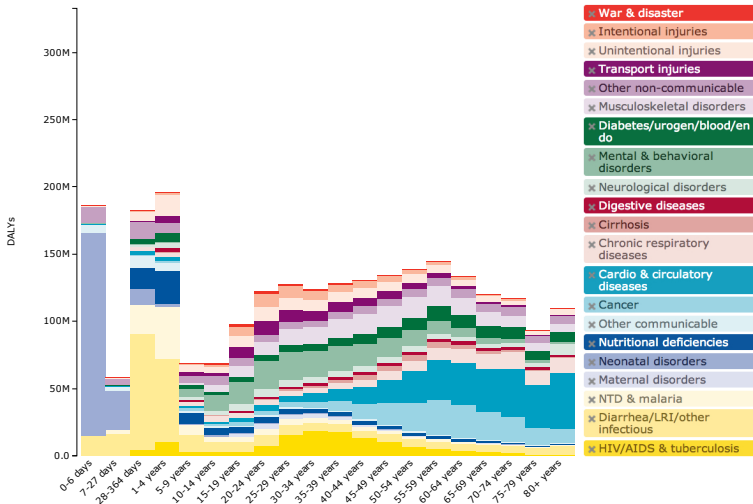
Years of Life Lost in 2010



Years Lost to Disability in 2010



Disability-Adjusted Life Years 2010



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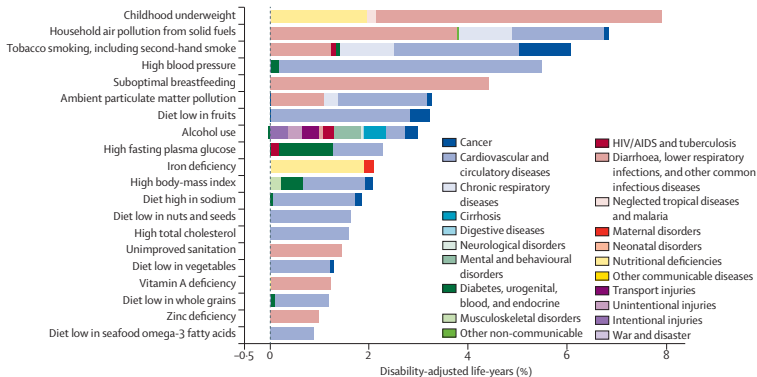
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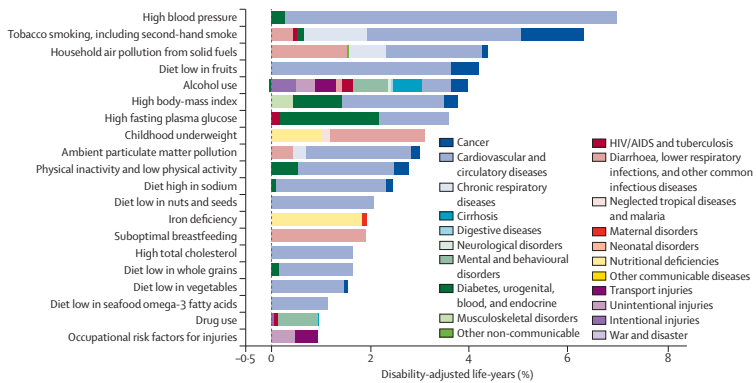
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Risk Factors in 1990

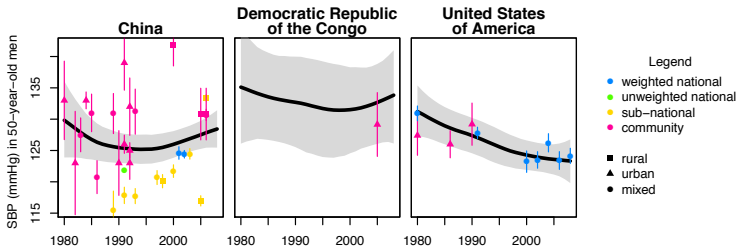


Risk Factors in 2010



Goal

Synthesize fragmentary data to make country-level estimates of time trends in risk factors by age and sex for all nations.

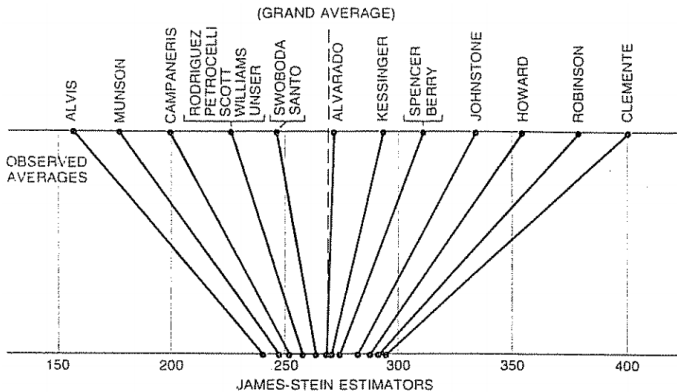


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- 1 Motivation: Estimating the Global Burden of Disease
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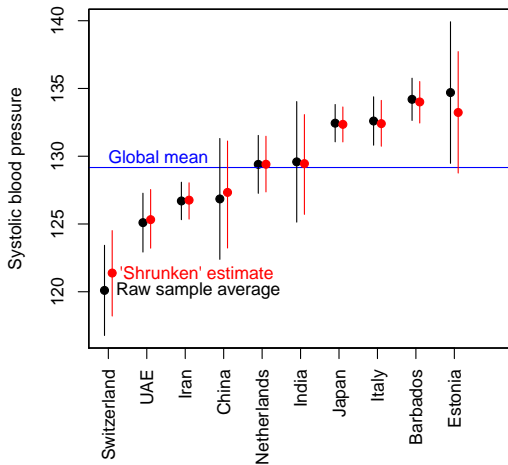
Stein's paradox:

'Shrink' raw data toward the mean

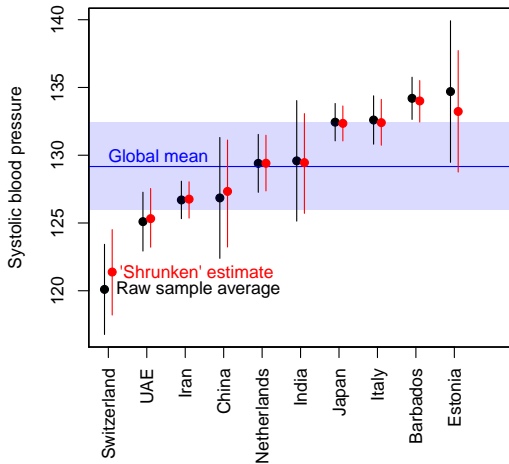


JAMES-STEIN ESTIMATORS for the 18 baseball players were calculated by “shrinking” the individual batting averages toward the overall “average of the averages.” In this case the grand average is .265 and each of the averages is shrunk about 80 percent of the distance to this value. Thus the theorem on which Stein’s method is based asserts that the true batting abilities are more tightly clustered than the preliminary batting averages would seem to suggest they are.

'Shrink' raw data toward the mean with a Bayesian model



'Shrink' raw data toward the mean with a Bayesian model



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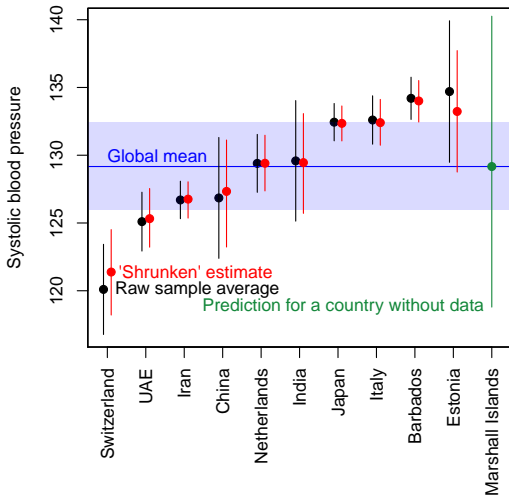
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Predict for countries without data with a Bayesian model

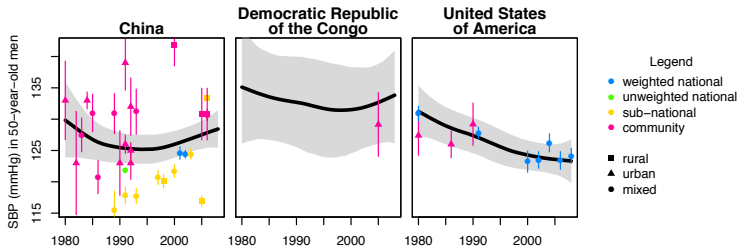


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Goal

Synthesize fragmentary data to make country-level estimates of time trends in risk factors by age and sex for all nations.



The likelihood

$$y_{h,i} \sim \mathcal{N} \left(\underbrace{a_{j[i]}^c + b_{j[i]}^c t_i}_{\text{geographic hierarchy}} + \underbrace{u_{j[i],t_i}}_{\text{nonlinear change in time}} + \underbrace{\gamma_i(z_h)}_{\text{flexible age model}} + \underbrace{X_i \beta}_{\text{covariate effects}} + \underbrace{e_i}_{\text{random effects}}, \underbrace{\frac{\overbrace{SD_{h,i}^2}}{\text{sampling var.}}}{n_{h,i}} + \underbrace{\tau_i^2}_{\text{residual var.}} \right)$$

Heteroscedastic random effects \Rightarrow discount unrepresentative studies

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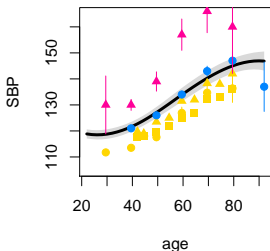
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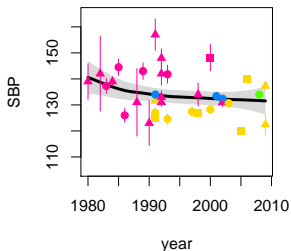
$$y_{h,i} \sim \mathcal{N} \left(\underbrace{a_{j[i]}^c + b_{j[i]}^c t_i}_{\text{geographic hierarchy}} + \underbrace{u_{j[i],t_i}}_{\text{nonlinear change in time}} + \underbrace{\gamma_i(z_h)}_{\text{flexible age model}} + \underbrace{X_i \beta}_{\text{covariate effects}} + \underbrace{e_i}_{\text{random effects}}, \underbrace{\frac{SD_{h,i}^2}{n_{h,i}}}_{\text{sampling var.}} + \underbrace{\tau_i^2}_{\text{residual var.}} \right)$$

Heteroscedastic random effects \Rightarrow discount unrepresentative studies

Chinese males, 1991



Chinese males, 55–64 y.



- weighted national
- unweighted national
- sub-national
- community
- rural
- ▲ urban
- mixed

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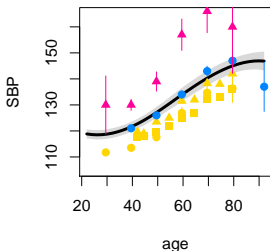
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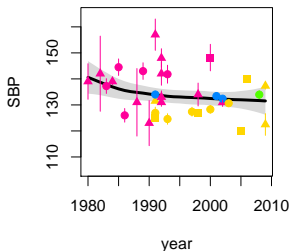
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Heteroscedastic random effects \Rightarrow discount unrepresentative studies

Chinese males, 1991



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$$\text{Var}(e_i) = \begin{cases} \nu_w & \text{if study } i \text{ is weighted national} \\ \nu_u & \text{if study } i \text{ is unweighted national} \\ \nu_s & \text{if study } i \text{ is sub-national} \\ \nu_c & \text{if study } i \text{ is community,} \end{cases}$$

with $\nu_w < \nu_u < \nu_s < \nu_c$.

Geographic hierarchy \Rightarrow shrinkage

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$$y_{h,i} \sim \mathcal{N} \left(\underbrace{a_{j[i]}^c + b_{j[i]}^c t_i}_{\text{geographic hierarchy}} + \underbrace{u_{j[i],t_i}}_{\text{nonlinear change in time}} + \underbrace{\gamma_i(z_h)}_{\text{flexible age model}} + \underbrace{X_i \beta}_{\text{covariate effects}} + \underbrace{e_i}_{\text{random effects}}, \underbrace{\frac{\overbrace{SD_{h,i}^2}}{\text{sampling var.}}}{n_{h,i}} + \underbrace{\tau_i^2}_{\text{residual var.}} \right)$$

Geographic hierarchy \Rightarrow shrinkage

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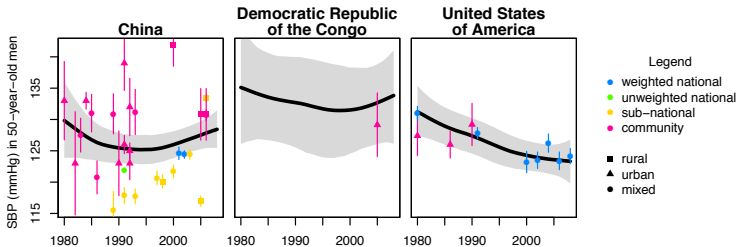
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Geographic hierarchy \Rightarrow shrinkage

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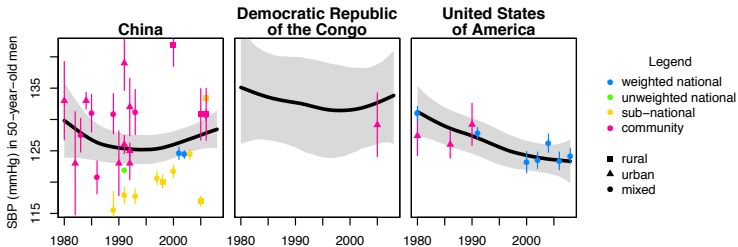
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$$\text{Country: } a_j^c \sim \mathcal{N}(a_{k[j]}^s, \kappa_a^c), \quad b_j^c \sim \mathcal{N}(b_{k[j]}^s, \kappa_b^c).$$

$$\text{Sub-region: } a_k^s \sim \mathcal{N}(a_{l[k]}^r, \kappa_a^s), \quad b_k^s \sim \mathcal{N}(b_{l[k]}^r, \kappa_b^s).$$

$$\text{Region: } a_l^r \sim \mathcal{N}(a^g, \kappa_a^r), \quad b_l^r \sim \mathcal{N}(b^g, \kappa_b^r).$$

Gaussian autoregressive priors \Rightarrow nonlinear change in time

$$y_{h,i} \sim \mathcal{N} \left(\underbrace{a_{j[i]}^c + b_{j[i]}^c t_i}_{\text{geographic hierarchy}} + \underbrace{u_{j[i],t_i}}_{\text{nonlinear change in time}} + \underbrace{\gamma_i(z_h)}_{\text{flexible age model}} + \underbrace{X_i \beta}_{\text{covariate effects}} + \underbrace{e_i}_{\text{random effects}}, \underbrace{\frac{SD_{h,i}^2}{n_{h,i}}}_{\text{sampling var.}} + \underbrace{\tau_i^2}_{\text{residual var.}} \right)$$

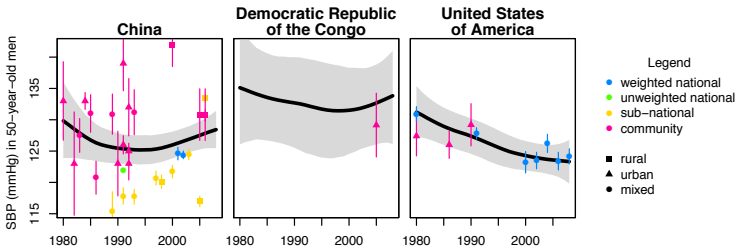
Gaussian autoregressive priors \Rightarrow nonlinear change in time

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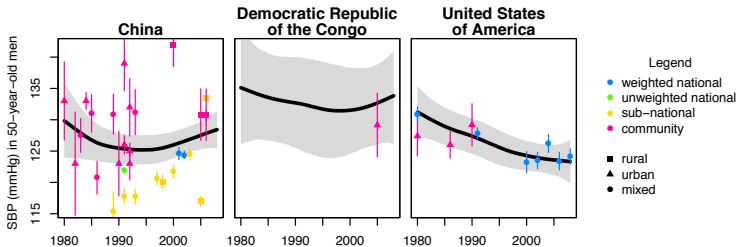
Gaussian autoregressive priors \Rightarrow nonlinear change in time

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$$u_j = u_j^c + u_{k[j]}^s + u_{l[k]}^r + u^g,$$

$$u_j^c \sim \mathcal{N}(0, (\lambda_c P)^{-1}) \quad \text{for } j = 1, \dots, J$$

$$u_k^s \sim \mathcal{N}(0, (\lambda_s P)^{-1}) \quad \text{for } k = 1, \dots, K$$

$$u_l^r \sim \mathcal{N}(0, (\lambda_r P)^{-1}) \quad \text{for } l = 1, \dots, L$$

$$u^g \sim \mathcal{N}(0, (\lambda_g P)^{-1}).$$

Cubic splines \Rightarrow flexible age model

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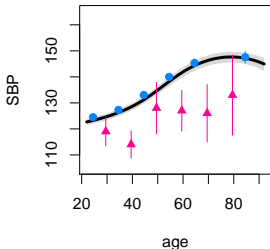
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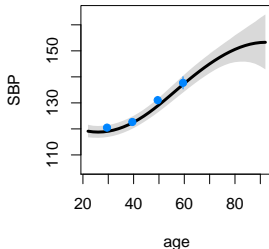
$$y_{h,i} \sim \mathcal{N} \left(\underbrace{a_{j[i]}^c + b_{j[i]}^c t_i}_{\text{geographic hierarchy}} + \underbrace{u_{j[i],t_i}}_{\text{nonlinear change in time}} + \underbrace{\gamma_i(z_h)}_{\text{flexible age model}} + \underbrace{X_i \beta}_{\text{covariate effects}} + \underbrace{e_i}_{\text{random effects}}, \underbrace{\frac{SD_{h,i}^2}{n_{h,i}}}_{\text{sampling var.}} + \underbrace{\tau_i^2}_{\text{residual var.}} \right)$$

Cubic splines \Rightarrow flexible age model

Japanese males, 1987



Singapore males, 1998



- weighted national
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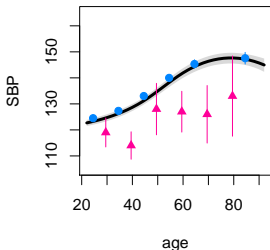
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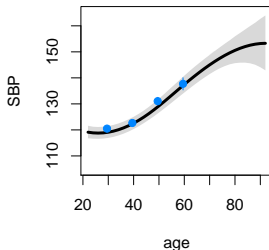
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Cubic splines \Rightarrow flexible age model

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$$\gamma_i(z_h) = \gamma_{1i}z_h + \gamma_{2i}z_h^2 + \gamma_{3i}z_h^3 + \gamma_{4i}(z_h - 45)_+^3 + \gamma_{5i}(z_h - 60)_+^3$$

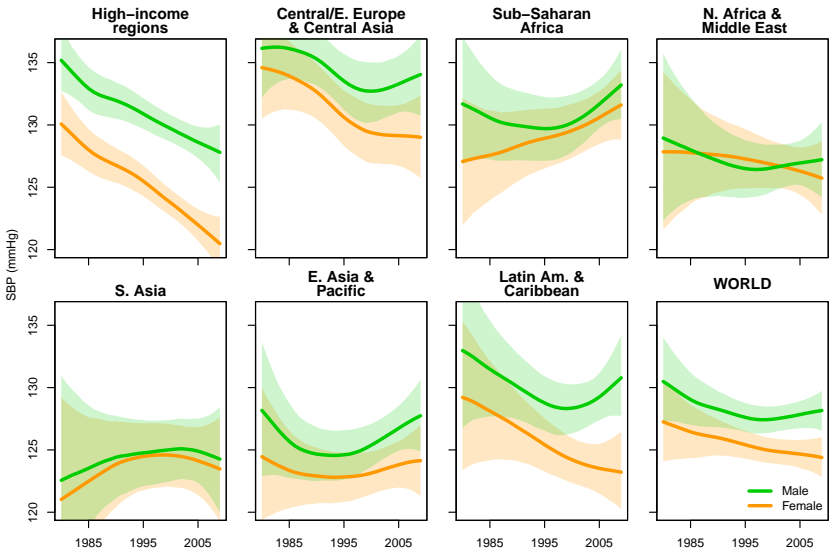
$$\gamma_{ki} = \psi_k + \phi_k\mu_i + c_{kj}[i], \quad k = 1, \dots, 5$$

$$\mu_i = a_{j[i]}^c + b_{j[i]}^c t_i + X_i\beta + u_{j[i], t_i}$$

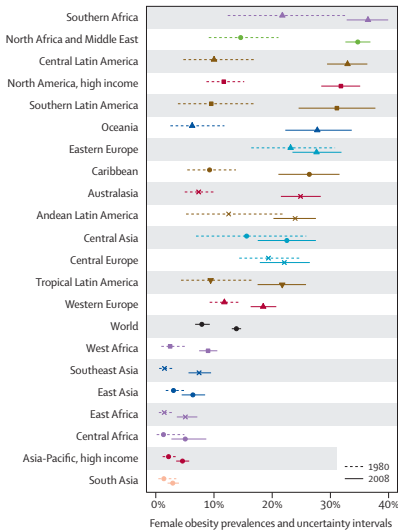
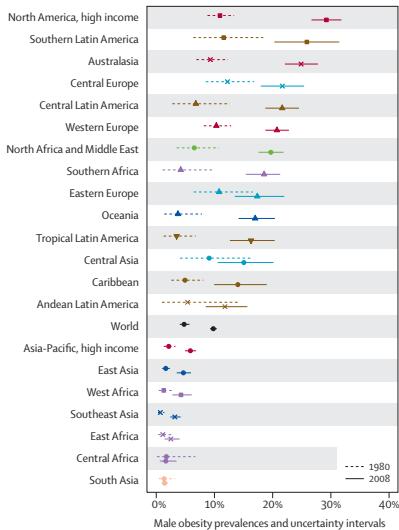
Covariates \Rightarrow better estimates when data are unavailable

$$y_{h,i} \sim \mathcal{N} \left(\underbrace{a_{j[i]}^c + b_{j[i]}^c t_i}_{\text{geographic hierarchy}} + \underbrace{u_{j[i],t_i}}_{\text{nonlinear change in time}} + \underbrace{\gamma_i(z_h)}_{\text{flexible age model}} + \underbrace{X_i \beta}_{\text{covariate effects}} + \underbrace{e_i}_{\text{random effects}}, \underbrace{\frac{\overbrace{SD_{h,i}^2}}{\text{sampling var.}}}_{\text{residual var.}} + \underbrace{\tau_i^2}_{\text{residual var.}} \right)$$

Inference



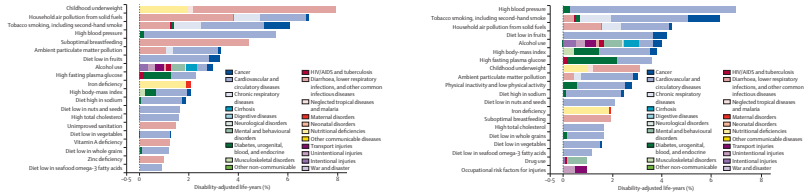
Other model applications: BMI

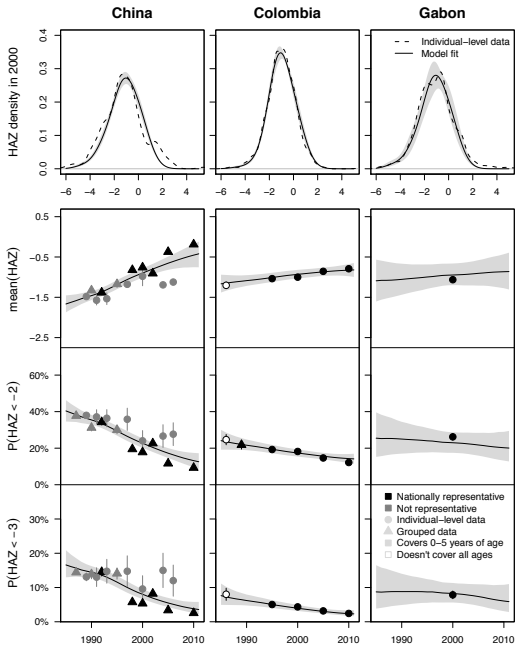


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The Global Burden of Disease: Risk Factors in 1990 and 2010





Goal

Synthesize
fragmentary data
to make
country-level
estimates of time
trends in **the full
distribution** of
risk factors for all
low- and
middle-income
nations.

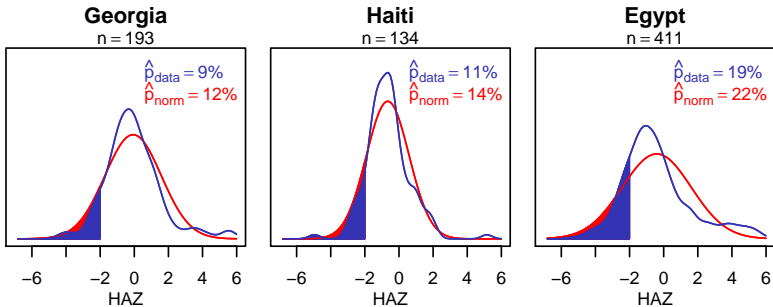
Height-for-age z-score distributions are skewed, and assuming normality \Rightarrow bias

Motivation

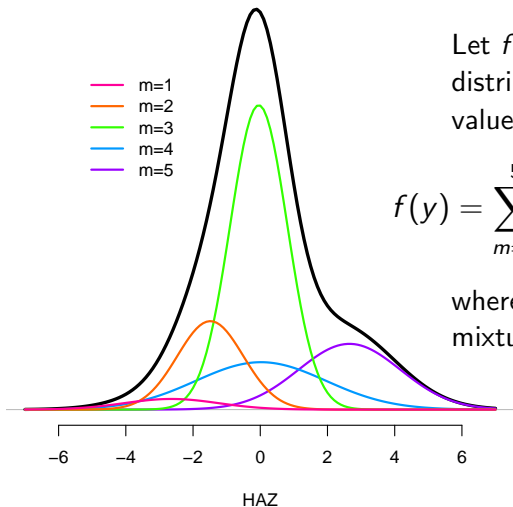
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Use a mixture of normals \Rightarrow estimate the shape of distributions



Let $f(y)$ be a
distribution of HAZ
values:

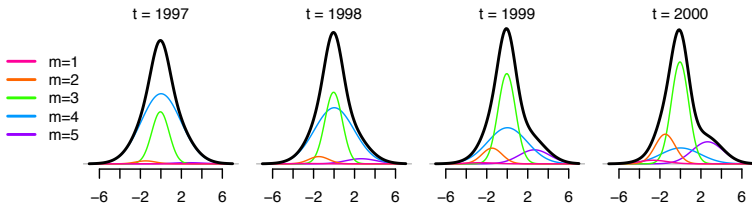
$$f(y) = \sum_{m=1}^5 w_m \mathcal{N}(\mu_m, \sigma_m^2),$$

where m denotes a
mixture component.

Allow the weights (w) to vary \Rightarrow a distribution for each study

Let $f_i(y)$ be the distribution of HAZ values in study i :

$$f_i(y) = \sum_{m=1}^5 w_{mi} \mathcal{N}(\mu_m, \sigma_m^2)$$

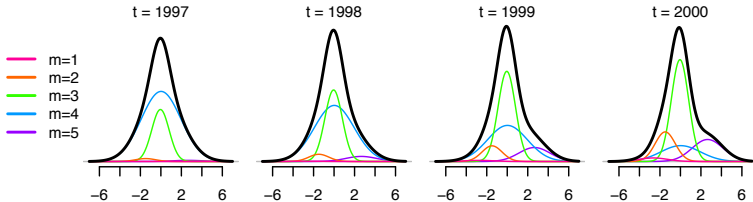


Borrow strength geographically, in time, and in covariates as in the SBP model

$$f_i(y) = \sum_{m=1}^5 w_{mi} \mathcal{N}(\mu_m, \sigma_m^2)$$

$$w_{mi} = \Phi(\alpha_{mi}) \prod_{u=1}^{m-1} (1 - \Phi(\alpha_{ui}))$$

$$\alpha_{mi} \sim \mathcal{N}(a_{mj[i]}^c + b_{mj[i]}^c t_i + u_{mj[i],t_i} + X_i \beta_m + e_{mi}, \tau_{mi}^2)$$



Combine individual-level data with summary statistics

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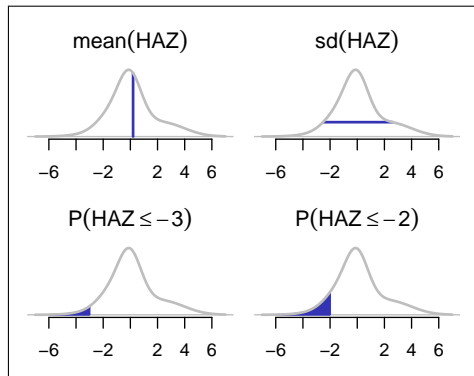
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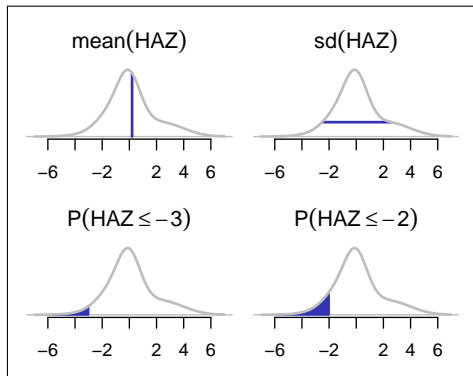
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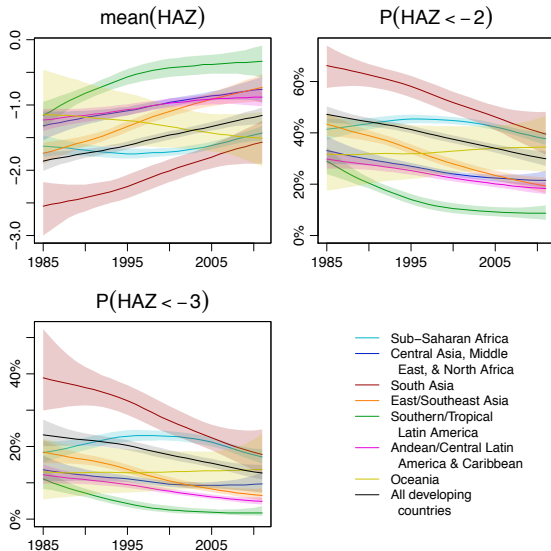


Combine individual-level data with summary statistics

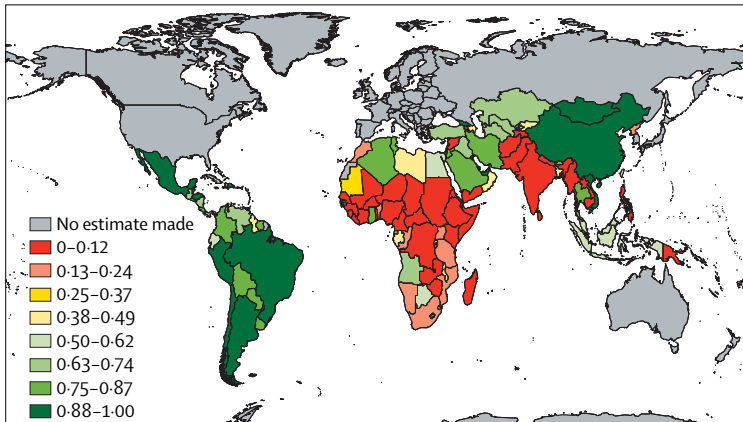


$$\begin{aligned} P(\theta|y) &\propto P(y|\theta) P(\theta) \\ &= P(y_{\text{individual}}|\theta) P(y_{\text{summary}}|\theta) P(\theta) \end{aligned}$$

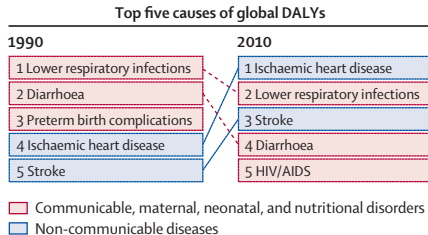
Inference



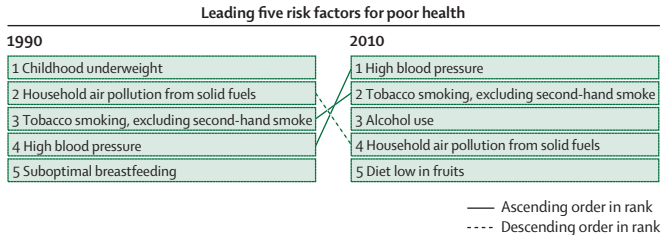
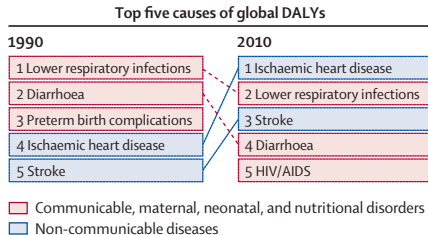
Other model applications: Posterior probability of meeting Millennium Development Goal #1



A global risk transition



A global risk transition



Why use a Bayesian approach for complex hierarchical analyses?

- To account for hyperparameter uncertainty
- To impose a penalty on model complexity
- To obtain, via MCMC:
 - Computational feasibility
 - Posterior draws for stakeholders
 - Inference on complex functionals

Thanks!

- Chris Paciorek (U.C. Berkeley)
- Majid Ezzati (Imperial College London)
- Gretchen Stevens (WHO)
- Goodarz Danaei (Harvard)

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Questions?

